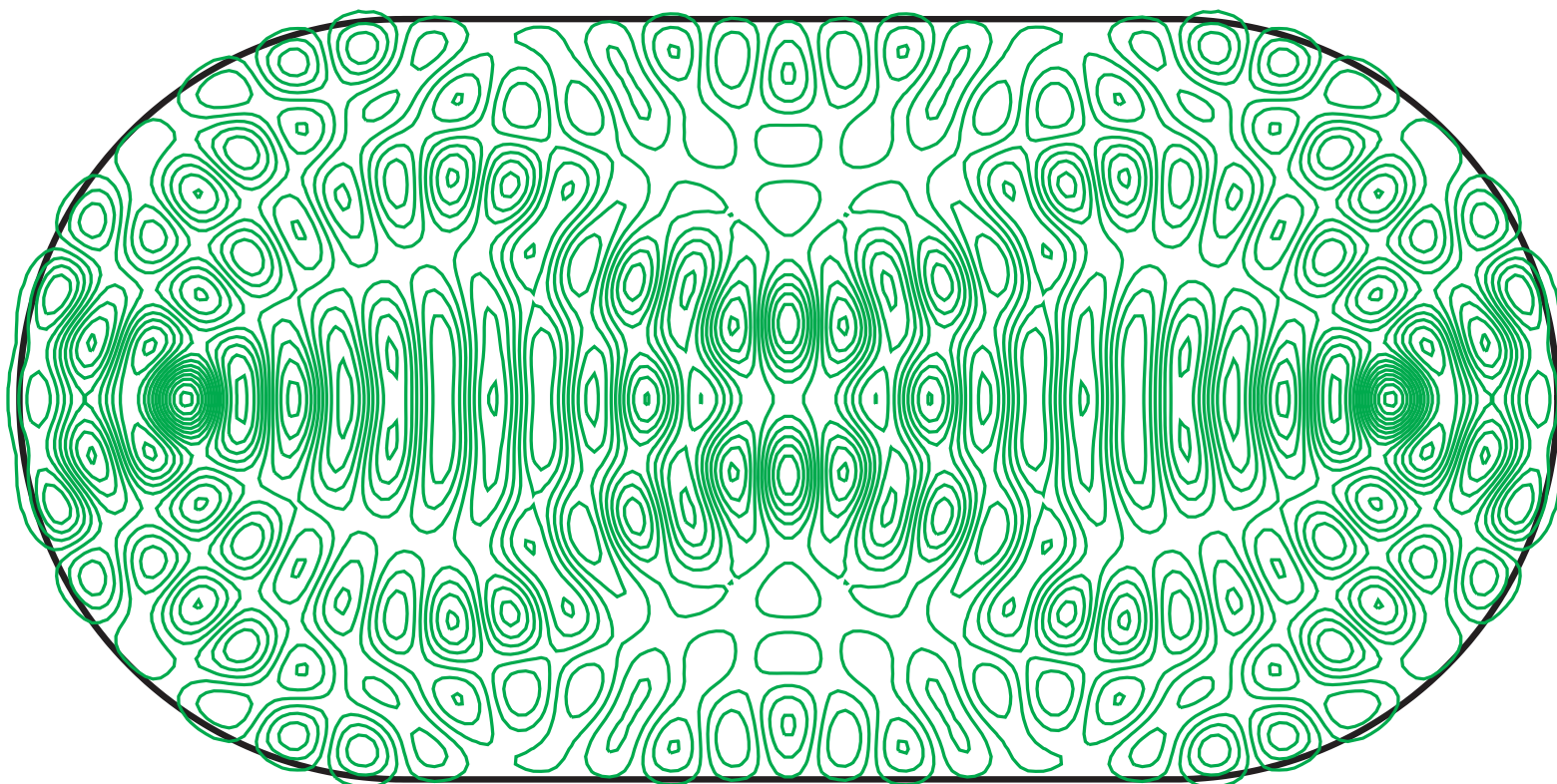


Extreme statistics of random and quantum chaotic states

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Today's Thread of Logic

- 1) The statistics of extreme eigenstate intensities
 - Densities and distribution functions
 - Largest and smallest intensities
 - Universal distributions
 - Weibull/Fréchet
 - Gumbel
- 2) Random states and chaotic quantum eigenstates
 - Complex random states and unitary ensembles¹
 - Exact results
 - Kicked rotor
 - Real random states and orthogonal ensembles
 - Saddle point approximations

¹recent published work: A. Lakshminarayan et al., *Phys. Rev. Lett.* **100**, 044103 (2008).

Distribution functions for maxima and minima intensities

- Suppose an ensemble of systems acts in an N -dimensional vector space, $\{|j\rangle\}$, $j = 1, \dots, N$ with eigenvectors of a member system, $\{|\phi_n\rangle\}$, $n = 1, \dots, N$. Then, the intensities for a single eigenstate are

$$s_j = |\langle \phi_n | j \rangle|^2$$

- Using n and the system ensemble, a joint intensity probability density can be defined and denoted

$$\rho(\vec{s}; N) = \rho(s_1, s_2, \dots, s_N; N)$$

- Let the maximum intensity be $s = \max [s_j]$, $j = 1, \dots, N$
- The distribution function is given by,

$$F_{max}(t; N) = \int_{\frac{1}{N}}^t ds \rho_{max}(s; N) = \int_0^t d\vec{s} \rho(\vec{s}; N) ; \left[\frac{d}{dt} F_{max}(t; N) = \rho_{max}(t; N) \right]$$

- Or for the minimum intensity $s = \min [s_j]$, $j = 1, \dots, N$

$$F_{min}(t; N) = 1 - \int_t^{\frac{1}{N}} ds \rho_{min}(s; N) = 1 - \int_t^1 d\vec{s} \rho(\vec{s}; N)$$

Distribution functions for maxima and minima uncorrelated variables

- Let the $\{s_j\}$, $j = 1, \dots, N$ be similarly distributed independent random variables.
- A joint probability density can be defined and denoted

$$\rho(\vec{s}; N) = \rho(s_1, s_2, \dots, s_N; N) = \prod_{j=1}^N \rho(s_j; N)$$

- The distribution function of the maximum is given by,

$$F_{max}(t; N) = \int^t d\vec{s} \rho(\vec{s}; N) = \left[\int^t ds_j \rho(s_j; N) \right]^N \quad (\text{any } j)$$

- Or for the minimum

$$F_{min}(t; N) = 1 - \int_t^{\infty} d\vec{s} \rho(\vec{s}; N) = 1 - \left[1 - \int_t^{\infty} ds_j \rho(s_j; N) \right]^N$$

Universal distribution functions - Fisher/Tippett 1928

uncorrelated variables (cont.)

- The Weibull/Fréchet distribution function

$$F(t; N) = 1 - \exp[-(\pm t - a_N)^{\gamma_N} / b_N]$$

is expected for uncorrelated random variables with compact support from above or below (or heavy tailed densities).

- The Gumbel distribution function

$$F(t; N) = \exp[-e^{-(t-a_N)/b_N}]$$

is expected for uncorrelated random variables with non-compact support whose tails decay at least exponentially fast.

- For example, consider the uniform density $\rho(t) = 1$ ($0 \leq t \leq 1$):

$$\begin{array}{llll} F_{max}(t; N) & = & t^N & \longrightarrow e^{-N(1-t)} & \text{Weibull} \\ F_{min}(t; N) & = & 1 - (1 - t)^N & \longrightarrow 1 - e^{-Nt} & \text{Fréchet} \end{array}$$

Two relevant examples: complex and real Gaussian amplitudes uncorrelated variables (cont.)

- Complex Gaussian amplitude leads to $\frac{1}{N}$ -mean intensity density:

$$\rho(t) = Ne^{-Nt} \quad (0 \leq t \leq \infty)$$

- and hence

$$F_{max}(t; N) = (1 - e^{-Nt})^N \rightarrow \exp\left(-e^{-N(t - \frac{1}{N} \ln N)}\right) \quad \text{Gumbel}$$

$$F_{min}(t; N) = 1 - [1 - (1 - e^{-Nt})]^N = 1 - e^{-N^2 t} \quad \text{Fréchet}$$

- Real Gaussian amplitude leads to $\frac{1}{N}$ -mean intensity density:

$$\rho(t) = \sqrt{\frac{N}{2\pi t}} e^{-Nt/2} \quad (0 \leq t \leq \infty)$$

- and hence

$$F_{max}(t; N) = \text{erf}^N\left(\sqrt{Nt/2}\right) \rightarrow \exp\left(-e^{-\frac{N}{2}\left(t - \frac{1}{N} \ln \frac{2N}{\pi t}\right)}\right) \quad \text{Gumbel?}$$

$$F_{min}(t; N) = 1 - \left[1 - \text{erf}\left(\sqrt{Nt/2}\right)\right]^N \rightarrow 1 - e^{-\sqrt{\frac{2N^3 t}{\pi}}} \quad \text{Fréchet}$$

Joint probability densities for intensities correlated variables

- A norm constraint is naturally expressed in amplitude variables:

$$\rho_{\beta}(z_1, z_2, \dots, z_N) = \frac{\Gamma\left(\frac{N\beta}{2}\right)}{\pi^{N\beta/2}} \delta\left(\sum_{j=1}^N |z_j|^2 - 1\right)$$

where $\beta = 1, 2$ for real and complex respectively. The real case corresponds to the orthogonal random matrix ensembles and the complex case to the unitary ensembles. Switching to intensities:

$$\rho_{\beta}(\vec{s}; N) = \pi^{N(\beta/2-1)} \Gamma\left(\frac{N\beta}{2}\right) \left[\prod_{j=1}^N s_j^{\beta/2-1} ds_j \right] \delta\left(\sum_{j=1}^N s_j - 1\right)$$

- The complex case is equivalent to the “broken stick problem” in which $N - 1$ cuts at uniformly random locations are made in a unit length stick.
- The real case is intimately connected to the relationship between hyperspherical and cartesian coordinates.

An auxiliary function for “decorrelating” intensities

- The distribution function for the maximum is:

$$F_{max}^{\beta}(t; N) = \pi^{N(\beta/2-1)} \Gamma\left(\frac{N\beta}{2}\right) \left[\prod_{j=1}^N \int_0^t s_j^{\beta/2-1} ds_j \right] \delta\left(\sum_{j=1}^N s_j - 1\right)$$

- Define the auxiliary function $G^{\beta}(t, u; N)$ which results from replacing unity in the norm constraint by u and thus, $F_{max}^{\beta}(t; N) = G^{\beta}(t, u = 1; N)$.
- The Laplace transform of $G^{\beta}(t, u; N)$ renders the integrals over the N differentials ds_j into a product form and gives:

$$\int_0^{\infty} e^{-us} G_{\beta}(t, N, u) du = \begin{cases} \Gamma\left(\frac{N}{2}\right) \left(\frac{\text{erf}(\sqrt{st})}{\sqrt{s}}\right)^N & \text{real} \\ \Gamma(N) \left(\frac{1-e^{-st}}{s}\right)^N & \text{complex} \end{cases}$$

- The N integrals have been performed at the cost of now needing the inverse Laplace transforms of these expressions.

Exact results for unitary ensembles

- The distribution function for the maximum follows by expanding the N^{th} power and using the inverse Laplace transform:

$$\mathcal{L}_s^{-1} \left(\frac{e^{-smt}}{s^N} \right) = \frac{1}{\Gamma(N)} (u - mt)^{N-1} \Theta(u - mt)$$

and gives

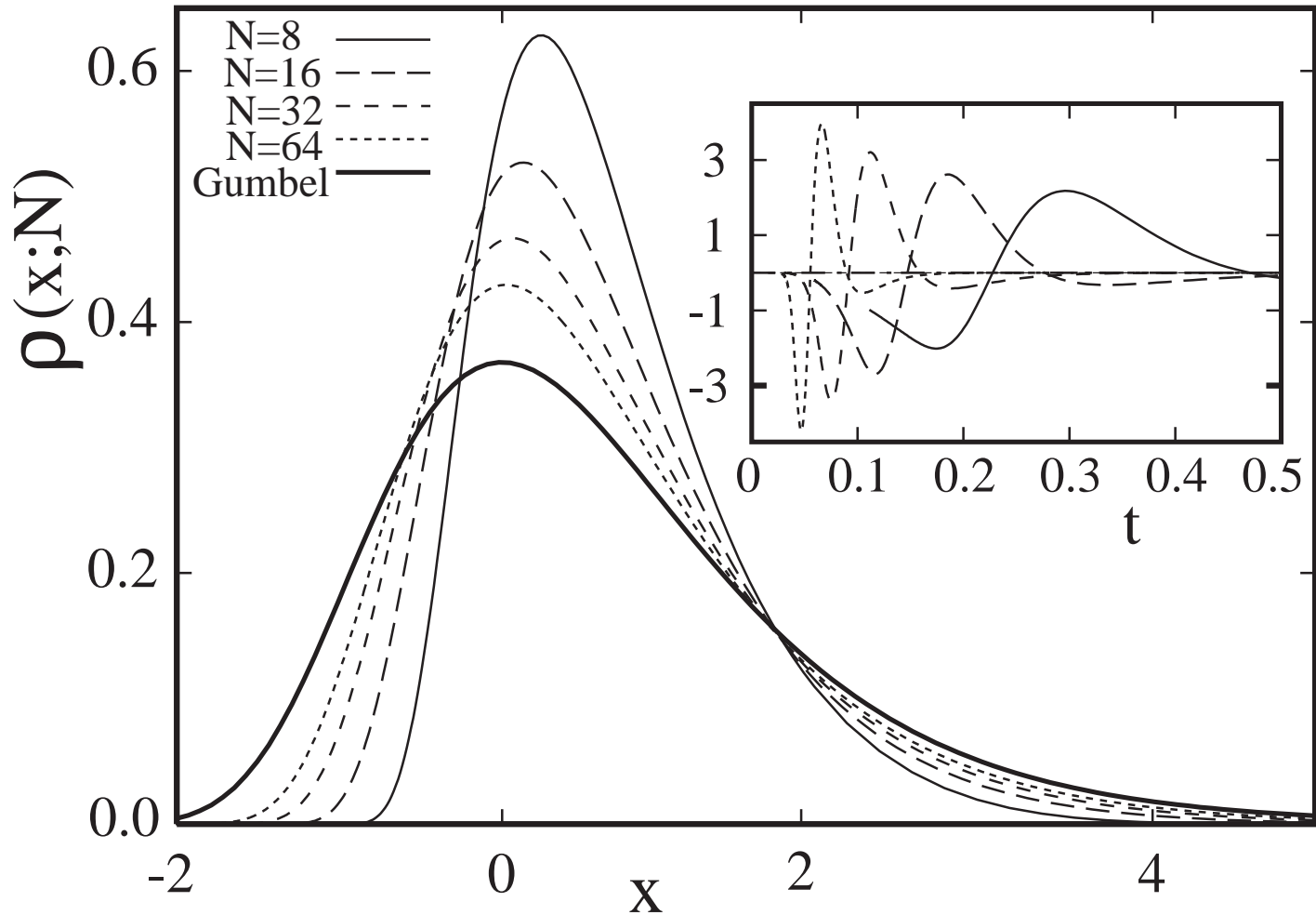
$$F_{max}^{\beta=2}(t; N) = \sum_{m=0}^N \binom{N}{m} (-1)^m (1 - mt)^{N-1} \Theta(1 - mt)$$

- Interestingly, this reduces to a piecewise smooth expression with the intervals $I_k = [1/(k+1), 1/k]$, where $k = 1, 2, \dots, N-1$

$$F_{max}(t \in I_k; N) = \sum_{m=0}^k \binom{N}{m} (-1)^m (1 - mt)^{N-1}$$

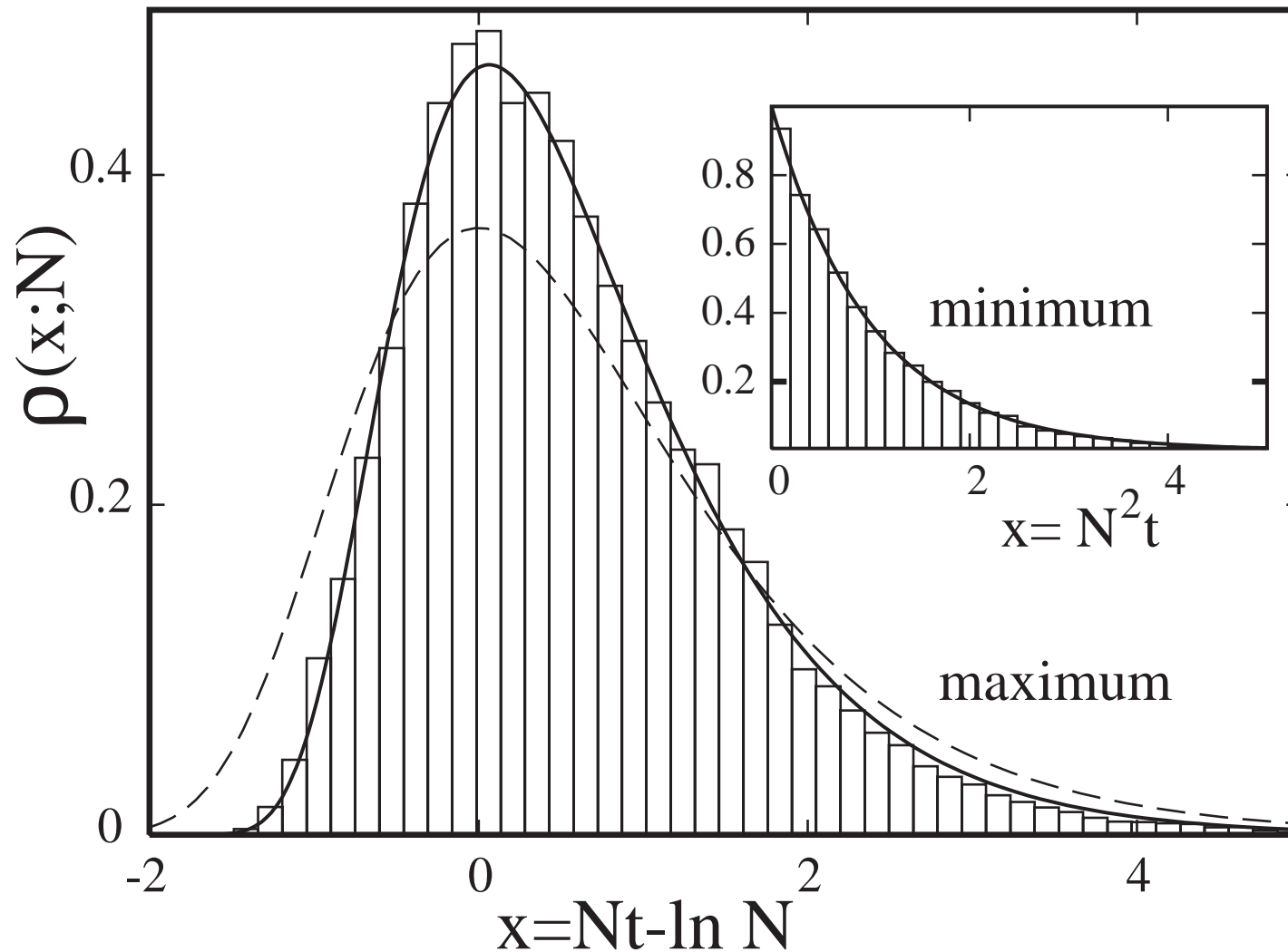
- All distributions resulting from correlated variables possessing at least unit norm constraints, satisfy a combinatoric form of this type.

Exact results for unitary ensembles (cont.)



The exact probability density and the asymptotic Gumbel density using the scaled variable $x = N(t - \ln(N))/N$ with increasing N . The inset shows the difference between the exact and the Gumbel densities for the same values of N , but in the unscaled variable.

The quantum kicked rotor as an example of a chaotic system



The probability densities (histograms) of the scaled maximum and minimum (inset) intensity of eigenfunctions in the position basis of the quantum kicked rotor for $N = 32$ in the parameter range $13.8 < K < 14.8$. Shown as a continuous line is the exact density for random states while the dotted ones are the respective Gumbel and Fréchet densities.

A saddle point approximation for the orthogonal ensembles

A saddle point approximation for the inverse Laplace transform:

$$\begin{aligned} & \mathcal{L}_s^{-1} \left((\pi st)^{m/2} e^{mst} \operatorname{erfc}^m(\sqrt{st}) \left(\frac{1}{\pi t}\right)^{m/2} \frac{e^{-smt}}{s^{\frac{N+m}{2}}} \right) \approx \left(\frac{N+m}{2-2mt}\right)^{m/2} \\ & \times \exp\left(\frac{m(N+m)t}{2-2mt}\right) \operatorname{erfc}^m\left(\sqrt{\frac{(N+m)t}{2-2mt}}\right) \frac{1}{\Gamma\left(\frac{N+m}{2}\right)} (1-mt)^{\frac{N+m}{2}-1} \Theta(1-mt) \end{aligned}$$

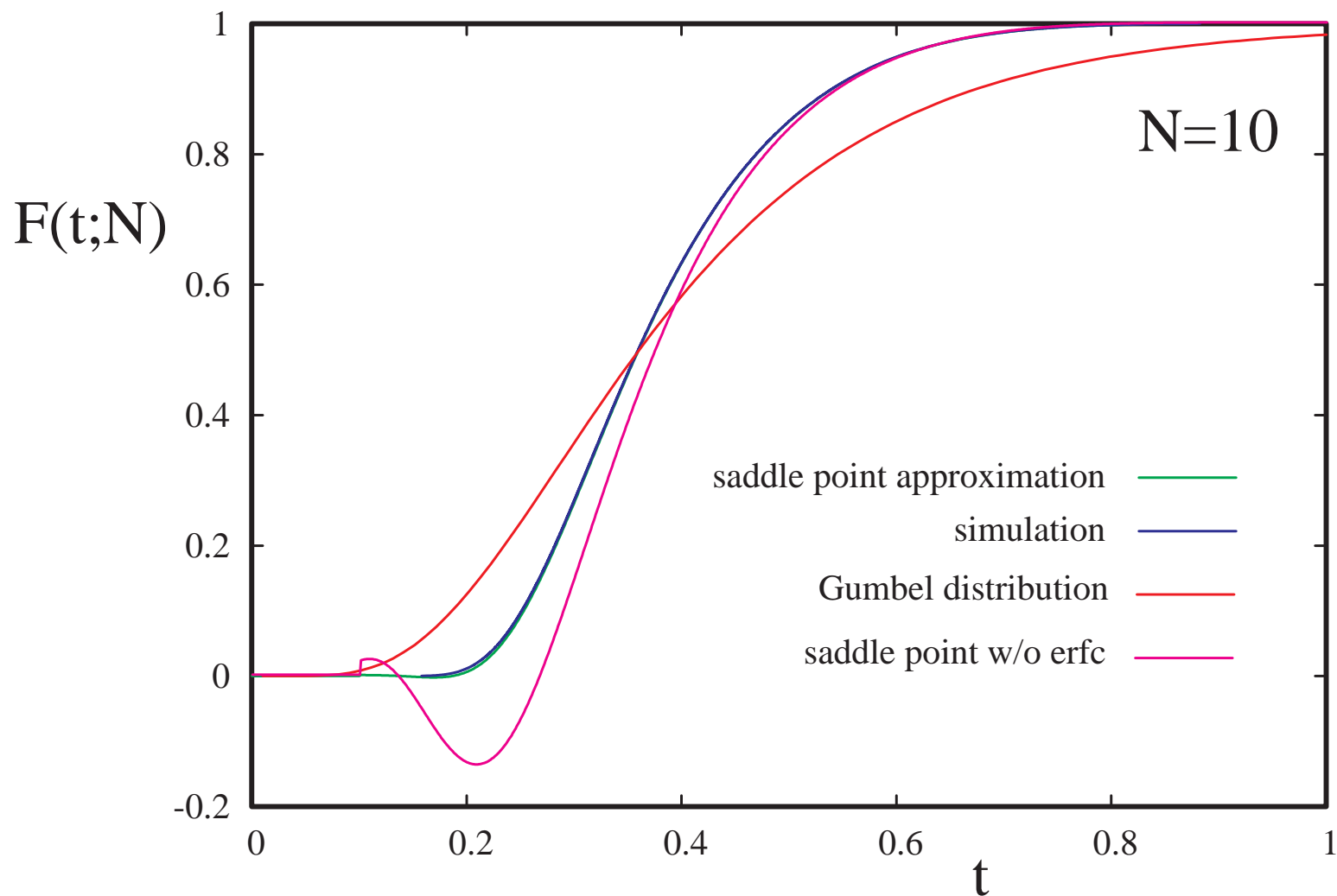
gives $F_{max}^{\beta=1}(t; N) =$

$$\sum_{m=0}^k \frac{\binom{N}{m} (-1)^m \left(\frac{N+m}{2}\right)^{\frac{m}{2}} \Gamma\left(\frac{N}{2}\right) (1-mt)^{\frac{N}{2}-1}}{\Gamma\left(\frac{N+m}{2}\right)} \exp\left(\frac{m\left(\frac{N+m}{2}\right)t}{1-mt}\right) \operatorname{erfc}^m\left(\sqrt{\frac{\left(\frac{N+m}{2}\right)t}{1-mt}}\right)$$

or more simply using only the asymptotic form of $\operatorname{erfc}(z)$,

$$F_{max}^{\beta=1}(t; N) = \sum_{m=0}^k \binom{N}{m} (-1)^m \frac{\Gamma\left(\frac{N}{2}\right)}{\Gamma\left(\frac{N+m}{2}\right)} \frac{(1-mt)^{\frac{N+m}{2}-1}}{(\pi t)^{\frac{m}{2}}}$$

The orthogonal ensembles



Comparison of the maximum intensity distribution functions for the orthogonal ensembles. The saddle point approximation improves considerably the agreement with the “exact” result from simulation vis-a-vis the asymptotic Gumbel form. The simpler form without the complementary error function is an improvement, but not good enough for small N to warrant its use.

Concluding remarks

- The statistical properties of extreme intensities has not previously been applied to understanding better the eigenstates of quantum systems.
- It is possible to derive some compact, exact results for the unitary ensembles with any dimensionality and give excellent approximations for the same quantities in the orthogonal ensemble.
- The maximum intensities tend to the infinite dimensional limit very slowly and thus the functional forms contain some information about system size. They further tend toward the Gumbel distribution although, a priori, one might have expected Weibull. Means scale as $\ln a_\beta N/N$ for unitary and (roughly for) orthogonal ensembles.
- The minimum intensity statistics tend much more rapidly toward their infinite dimensional Fréchet limiting form. The mean minima are N^{-2} and πN^{-3} for the unitary and orthogonal ensembles respectively.
- These “extreme” measures give us a new way to explore non-ergodic eigenstate behaviors.
 - It would be worthwhile exploring: i) other measures such as the intensity densities in the neighborhood of the maxima or minima, and ii) how system dynamics lead to deviations from the chaotic statistical extremes.